

An Improved Regression Type Estimator of Finite Population Mean using Coefficient of variation in PPS sampling.

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Abstract: Coefficient of variation is a unitless measure of dispersion and is widely used rather than standard deviation. In this paper for estimating population mean, auxiliary information is used in the form of coefficient of variation. The estimator depending on the estimated optimum value of k constant based on sample value under pps sampling scheme have been proposed. Its Bias, MSE and comparative study with the usual ratio, product and linear regression estimators is made. An empirical study is also included for the practical utility of the proposed estimator.

Key words :- Coefficient of variation Regression ,Ratio ,Product Estimators Bias & Mean Square Error Efficiency

I. Introduction

It is well know that at large scale surveys use of multiple auxiliary characteristics improves the precision of the estimators. In many practical situations, the coefficient of variation Cx of auxiliary variable x may be available. Auxiliary variable x in the form of known population mean are used in estimators like ratio, product and regression in the literature. We can use auxiliary variable in the form of coefficient of variation for the efficient estimation of population mean of study variable y. In this paper two auxiliary characters have been used in different ways as one for the selection of sample and other for the purpose of estimation to estimate the population mean.

Notations

Let y_i, x_{1i} and x_{2i} be two variables of characters y under study and values of the two auxiliary characters x_1 and x_2 for the i^{th} unit in the population ($i=1, 2, 3, \dots, N$) of size N. Let a sample of size n be drawn with ppswr sampling (based on x_1) and Y, X_1, X_2 be population total of y, x_1, x_2 respectively.

$$u_i = \frac{y_i}{Np_{1i}}, v_i = \frac{x_{2i}}{Np_{1i}}, p_{1i} = \frac{X_{1i}}{X_1}$$

$$\bar{u}_n = \frac{1}{n} \sum_{i=1}^n u_i = \bar{y}_{1pps}$$

$$\bar{v}_n = \frac{1}{n} \sum_{i=1}^n v_i = \bar{x}_{2pps}$$

$$S_u^2 = \sum_{i=1}^N p_{1i} (u_i - \bar{y})^2, s_u^2 = \sum_{i=1}^N p_{1i} (u_i - \bar{u}_n)^2$$

$$S_v^2 = \sum_{i=1}^N p_{1i} (v_i - \bar{x}_2)^2, s_v^2 = \sum_{i=1}^N p_{1i} (v_i - \bar{v}_n)^2$$

$$S_{uv} = \sum_{i=1}^N p_{1i} (u_i - \bar{y})(v_i - \bar{x}_2), b = \frac{S_{uv}}{S_v^2}$$

$$\mu_{rs} = \sum_{i=1}^N p_{1i} (u_i - \bar{y})^r (v_i - \bar{x}_2)^s$$

$$C_u = \frac{\sqrt{\mu_{20}}}{\bar{Y}}, C_v = \frac{\sqrt{\mu_{02}}}{\bar{X}_2}, \beta = \frac{S_{uv}}{S_v^2} = \rho \frac{S_u}{S_v}$$

$$\rho = \frac{\mu_{11}}{\sqrt{\mu_{20}\mu_{02}}} = \frac{S_{uv}}{S_u S_v}$$

$$\bar{y}_R = \frac{\bar{u}_n}{\bar{v}_n} \bar{X}_2 \text{-----(1.1)}$$

$$\bar{y}_p = \frac{\bar{u}_n \bar{v}_n}{\bar{X}_2} \text{-----(1.2)}$$

$$\bar{y}_h = \bar{y} + b(\bar{X} - \bar{x}) \text{-----(1.3)}$$

$$\beta_1(v) = \frac{\mu_{40}}{\mu_{20}^2} = \frac{\mu_{40}}{\bar{Y}^4 C_u^4}$$

$$\gamma_1(v) = \frac{\mu_{03}}{\mu_{02}^{3/2}} = \frac{\mu_{03}}{\bar{X}_2^3 C_v^3}$$

II. Proposed Estimator

$$\bar{y}_{Ipps} = \bar{y} + b(\bar{X} - \bar{x}) + k(\bar{x} - \frac{S_x}{c_x}) \text{-----(2.1)}$$

Bias and Mean Square Error of \bar{y}_{Ipps}

Let

$$\bar{y} = \bar{Y} + e_0, \bar{x} = \bar{X} + e_1$$

$$S_{uv} = S_{uv} + e_2, S_v^2 = S_v^2 + e_3$$

So that

$$E(e_0) = E(e_1) = E(e_2) = E(e_3) = 0$$

$$\text{and } E(e_0^2) = \frac{S_u^2}{n}, E(e_1^2) = \frac{S_v^2}{n}, E(e_3^2) = \beta_1(v) - 1 \frac{S_v^4}{n},$$

$$E(e_0 e_1) = \frac{S_{uv}}{n} = \frac{\beta S_v^2}{n}, E(e_0 e_3) = \frac{\mu_{12}}{n}, E(e_1 e_2) = \frac{\mu_{12}}{n}, E(e_1 e_3) = \frac{\mu_{03}}{n}$$

From (2.1) we have

$$\bar{y}_{Ipps} = (\bar{Y} + e_0) + (-e_1) \left(\frac{S_{uv} + e_2}{S_v^2 + e_3} \right) + k[\bar{X} + e_1 - \bar{X}_2 \left(\frac{S_v^2 + e_3}{S_v} \right)^{1/2}]$$

$$(\bar{y}_{Ipps} - \bar{Y}) = e_0 - \beta e_1 \left(1 + \frac{e_2}{S_{uv}} \right) \left(1 + \frac{e_3}{S_v^2} \right)^{-1} + k[\bar{X} + e_1 - \bar{X}_2 \left(1 + \frac{e_3}{S_v^2} \right)^{1/2}]$$

Expanding $(1 + \frac{e_3}{S_v^2})^{-1}$, $(1 + \frac{e_3}{S_v^2})^{1/2}$ and multiplying out and neglecting the terms of e_i^3 of higher degree we obtain.

$$(\bar{y}_{Ipps} - \bar{Y}) = (e_0 - \beta e_1) + k(\bar{X} + e_1) - k\bar{X}_2 \left(1 + \frac{e_3}{2S_v^2} - \frac{e_3^2}{8S_v^4} \right) + \beta \left(\frac{e_1 e_3}{S_v^2} - \frac{e_1 e_2}{S_{uv}} \right)$$

$$(\bar{y}_{Ipps} - \bar{Y}) = (e_0 - \beta e_1) + k(\bar{X} - \bar{X}_2) + k e_1 - k\bar{X}_2 \left(\frac{e_3}{2S_v^2} \right) + \beta \left(\frac{e_1 e_3}{S_v^2} - \frac{e_1 e_2}{S_{uv}} \right) + \frac{k\bar{X}_2 e_3^2}{8S_v^4} + \dots \text{-----(2.2)}$$

taking expectation on both sides of (2.2) we have the bias of (\bar{y}_{Ipps}) upto higher order

$$\text{Bias}(\bar{y}_{Ipps}) = \frac{\beta}{n} \left(\frac{\mu_{03}}{S_v^2} - \frac{\mu_{12}}{S_{uv}} \right) + \frac{k\bar{X}_2}{8n} (\beta_1(v) - 1)$$

$$= \frac{1}{n} \left[\left(\beta \frac{\mu_{03}}{S_v^2} - \frac{\mu_{12}}{S_{uv}} \right) + \frac{k\bar{X}_2}{8} (\beta_1(v) - 1) \right] \text{-----(2.3)}$$

Squaring both sides of (2.2) and taking expectation upto terms of higher order is

$$\begin{aligned}
 \text{MSE}(\bar{Y}_{Ipps}) &= (e_0 - \beta e_1)^2 + k^2 \left(e_1 - \frac{\bar{X}_2 e_3}{2S_v^2} \right)^2 + \beta^2 \left(\frac{e_1 e_3}{S_v^2} - \frac{e_1 e_2}{S_{uv}} \right)^2 + \frac{k^2 \bar{X}_2^2 e_3^4}{64S_v^4} + 2k(e_0 - \beta e_1) \left(e_1 - \frac{\bar{X}_2 e_3}{2S_v^2} \right) \\
 \text{MSE}(\bar{Y}_{Ipps}) &= E(e_0)^2 + \beta^2 E(e_1)^2 - 2\beta E(e_0 e_1) + k^2 [E(e_1)^2 + E\left(\frac{\bar{X}_2^2 e_3^2}{4S_v^4}\right) - \frac{\bar{X}_2 E(e_1 e_3)}{S_v^2}] + 2k(e_0 - \beta e_1) \left(e_1 - \frac{\bar{X}_2 e_3}{2S_v^2} \right) \\
 &= E(e_0)^2 + \beta^2 E(e_1)^2 - 2\beta E(e_0 e_1) + k^2 [E(e_1)^2 + E\left(\frac{\bar{X}_2^2 e_3^2}{4S_v^4}\right) - \frac{\bar{X}_2 E(e_1 e_3)}{S_v^2}] + 2kE\left[e_0 e_1 - \frac{\bar{X}_2 (e_0 e_3)}{2S_v^2} + \frac{\beta \bar{X}_2 (e_0 e_3)}{2S_v^2} - \beta (e_1^2)\right] \\
 &= \frac{\bar{Y}^2 C_u^2 (1-\rho^2)}{n} + k^2 \left[\frac{S_v^2}{n} + \frac{\bar{X}_2^2}{4n} \beta_1(v) - 1 - \frac{\bar{X}_2 \mu_{03}}{S_v^2 n} \right] + 2k \left[\frac{S_{uv}}{n} - \frac{\mu_{12}}{2n \bar{X}_2 C_v^2} + \frac{\beta \bar{X}_2 \mu_{03}}{2n S_v^2} - \beta \frac{S_v^2}{n} \right] \\
 &= \frac{\bar{Y}^2 C_u^2 (1-\rho^2)}{n} + \frac{k^2 \bar{X}_2^2}{4n} [4C_v^2 + \beta_1(v) - 1 - 4\gamma_1(v) C_v] + k \left[\frac{S_{uv} \gamma_1(v)}{n C_v} - \frac{\mu_{12}}{n \bar{X}_2 C_v^2} \right] \\
 &= \frac{\bar{Y}^2 C_u^2 (1-\rho^2)}{n} + \frac{k^2 \bar{X}_2^2}{4n} [4C_v^2 + \beta_1(v) - 1 - 4\gamma_1(v) C_v] + k \left[\frac{\bar{X}_2 C_v \gamma_1(v) S_{uv} - \mu_{12}}{n \bar{X}_2 C_v^2} \right] \dots (2.4)
 \end{aligned}$$

The optimum value of k & Minimum Mean Square Error of (\bar{Y}_{Ipps}) in (2.4) is given by

$$k_{opt} = \frac{2\bar{X}_2 C_v \gamma_1(v) S_{uv} - \mu_{12}}{\bar{X}_2^2 C_v^2 [4C_v^2 + \beta_1(v) - 1 - 4\gamma_1(v) C_v]} \dots (2.5)$$

$$\text{Minimum MSE}(\bar{Y}_{Ipps}) = \frac{\bar{Y}^2 C_u^2 (1-\rho^2)}{n} - \frac{[\bar{X}_2 C_v \gamma_1(v) S_{uv} - \mu_{12}]^2}{n \bar{X}_2^4 C_v^4 [4C_v^2 + \beta_1(v) - 1 - 4\gamma_1(v) C_v]} \dots (2.6)$$

III. Estimator Based on Estimated optimum K

The value of k may not be known for us practically so we replace by its estimated S_{uv} , μ_{12} , $\gamma_1(v)$, $\beta_1(v)$, μ_{03} & μ_{04} is involved in optimum K is

$$\hat{S}_{uv} = s_{uv} = \sum_{i=1}^n p_{ii} (u_i - \bar{u}_n)(v_i - \bar{v}_n)$$

$$\hat{\mu}_{12} = \sum_{i=1}^n p_{ii} (u_i - \bar{u}_n)(v_i - \bar{v}_n)^2$$

$$\hat{\gamma}_1(v) = \frac{\hat{\mu}_{03}}{\bar{X}_2^3 C_v^3}, \hat{\beta}_1(v) = \frac{\hat{\mu}_{04}}{\bar{X}_2^4 C_v^4}$$

$$\hat{\mu}_{03} = \sum_{i=1}^n p_{ii} (v_i - \bar{v}_n)^3$$

$$\hat{\mu}_{04} = \sum_{i=1}^n p_{ii} (v_i - \bar{v}_n)^4$$

we get the estimated optimum \hat{k} as

$$\begin{aligned}
 \hat{k} &= \frac{2[\bar{X}_2 C_v \hat{\gamma}_1(v) s_{uv} - \hat{\mu}_{12}]}{\bar{X}_2^3 C_v^2 [4C_v^2 + \hat{\beta}_1(v) - 1 - 4\hat{\gamma}_1(v) C_v]} \\
 \hat{k} &= \frac{2\left[\frac{\bar{X}_2 C_v \hat{\mu}_{03} s_{uv}}{\bar{X}_2^3 C_v^3} - \hat{\mu}_{12}\right]}{\bar{X}_2^3 C_v^2 \left[4C_v^2 + \frac{\hat{\mu}_{04}}{\bar{X}_2^4 C_v^4} - 1 - 4\frac{\hat{\mu}_{03} C_v}{\bar{X}_2^3 C_v^3}\right]} \dots (3.1)
 \end{aligned}$$

Estimator of population Mean \bar{Y} based on estimated optimum \hat{k} is

$$\bar{Y}_{Ipps} = \bar{y} + b(\bar{X} - \bar{x}) + \hat{k}(\bar{x} - \frac{S_x}{c_x}) \dots (3.2)$$

let $\hat{\mu}_{03} = \mu_{03} + e_4$, $\hat{\mu}_{12} = \mu_{12} + e_5$, $\hat{\mu}_{04} = \mu_{04} + e_6$

so that from (3.1), we get

$$\begin{aligned}
 \hat{k} &= \frac{2[\bar{X}_2 C_v (\mu_{03} + e_4)(s_{uv} + e_2) - \mu_{12} + e_5]}{\bar{X}_2^3 C_v^2 [4C_v^2 + \frac{\mu_{04} + e_6}{\bar{X}_2^4 C_v^4} - 1 - \frac{4(\mu_{03} + e_3) C_v}{\bar{X}_2^3 C_v^3}]} \\
 \hat{k} &= \frac{2\left[\frac{\bar{X}_2 C_v \mu_{03} s_{uv}}{\bar{X}_2^3 C_v^3} + \frac{\mu_{03} e_2 + s_{uv} e_4 + e_2 e_4}{\bar{X}_2^3 C_v^2} - \mu_{12} - e_5\right]}{\bar{X}_2^3 C_v^2 \left[4C_v^2 + \frac{\mu_{04}}{\bar{X}_2^4 C_v^4} + \frac{e_6}{\bar{X}_2^4 C_v^4} - 1 - \frac{4\mu_{03} C_v}{\bar{X}_2^3 C_v^3} - \frac{4e_3}{\bar{X}_2^3 C_v^2}\right]} \\
 &= \frac{2[\bar{X}_2 C_v \gamma_1(v) s_{uv} - \mu_{12} + \frac{\mu_{03} e_2 + s_{uv} e_4 + e_2 e_4 - \bar{X}_2^2 C_v^2 e_3}{\bar{X}_2^3 C_v^2}]}{\bar{X}_2^3 C_v^2 [4C_v^2 + \beta_1(v) - 1 - 4\gamma_1(v) C_v + \frac{e_6 - 4\bar{X}_2 C_v^2 e_4}{\bar{X}_2^4 C_v^4}]} \\
 &= \frac{2[\bar{X}_2 C_v \gamma_1(v) s_{uv} - \mu_{12}]}{\bar{X}_2^3 C_v^2 [4C_v^2 + \beta_1(v) - 1 - 4\gamma_1(v) C_v]} \left[1 + \frac{\mu_{03} e_2 + s_{uv} e_4 + e_2 e_4 - \bar{X}_2^2 C_v^2 e_3}{\bar{X}_2^3 C_v^2 [\bar{X}_2 C_v \gamma_1(v) s_{uv} - \mu_{12}]} \right] \left[1 + \frac{e_6 - 4\bar{X}_2 C_v^2 e_4}{\bar{X}_2^4 C_v^4 [4C_v^2 + \beta_1(v) - 1 - 4\gamma_1(v) C_v]} \right]^{-1} \\
 &= \frac{2[\bar{X}_2 C_v \gamma_1(v) s_{uv} - \mu_{12}]}{\bar{X}_2^3 C_v^2 [4C_v^2 + \beta_1(v) - 1 - 4\gamma_1(v) C_v]} \left[1 + \frac{\mu_{03} e_2 + s_{uv} e_4 + e_2 e_4 - \bar{X}_2^2 C_v^2 e_3}{\bar{X}_2^3 C_v^2 [\bar{X}_2 C_v \gamma_1(v) s_{uv} - \mu_{12}]} - \frac{e_6 - 4\bar{X}_2 C_v^2 e_4}{\bar{X}_2^4 C_v^4 [4C_v^2 + \beta_1(v) - 1 - 4\gamma_1(v) C_v]} + \dots \right] \dots (3.3)
 \end{aligned}$$

from (3.2) we have

$$(\bar{Y}_{1pps} - \bar{Y}) = (e_0 - \beta e_1) + k(\bar{X} - \bar{X}_2) + k(e_1 - \bar{X}_2 \frac{e_2}{2S_v^2}) + \beta(\frac{e_1 e_2}{S_v^2} - \frac{e_1 e_2}{S_{uv}}) + \frac{kX_2 e_2^2}{8S_v^4} + \dots \dots \dots (3.4)$$

Substituting k from (3.3) in (3.4), Squaring both sides and taking expectation, the MSE of \bar{Y}_{1pps} to the terms of higher order is

$$\begin{aligned} \text{MSE}(\bar{Y}_{1pps}) &= E[(e_0 - \beta e_1) - \frac{2[\bar{X}_2 C_v \gamma_1(v) S_{uv} - \mu_{12}]}{\bar{X}_2^2 C_v^2 [4C_v^2 + \beta_1(v) - 1 - 4\gamma_1(v) C_v]} (e_1 - \bar{X}_2 \frac{e_2}{2S_v^2})]^2 \\ &= \frac{\bar{Y}^2 C_u^2 (1 - \rho^2)}{n} + \frac{[\bar{X}_2 C_v \gamma_1(v) S_{uv} - \mu_{12}]^2}{n \bar{X}_2^4 [4C_v^2 + \beta_1(v) - 1 - 4\gamma_1(v) C_v]} - \frac{2[\bar{X}_2 C_v \gamma_1(v) S_{uv} - \mu_{12}]^2}{n \bar{X}_2^4 C_v^2 [4C_v^2 + \beta_1(v) - 1 - 4\gamma_1(v) C_v]} \\ &= \frac{\bar{Y}^2 C_u^2 (1 - \rho^2)}{n} - \frac{[\bar{X}_2 C_v \gamma_1(v) S_{uv} - \mu_{12}]^2}{n \bar{X}_2^4 [4C_v^2 + \beta_1(v) - 1 - 4\gamma_1(v) C_v]} \dots \dots \dots (3.5) \end{aligned}$$

Which shows that the mean square error of the estimator \bar{Y}_{1pps} in (3.2) based on the estimated optimum to the terms of higher order is same as \bar{Y}_{1pps} in (2.6).

IV. Comparison of the proposed Estimator

4.11 $\text{MSE}(\bar{Y}_{1pps}) < \text{MSE}(\bar{Y})$

$$\begin{aligned} \frac{\bar{Y}^2 C_u^2 (1 - \rho^2)}{n} - \frac{[\bar{X}_2 C_v \gamma_1(v) S_{uv} - \mu_{12}]^2}{n \bar{X}_2^4 [4C_v^2 + \beta_1(v) - 1 - 4\gamma_1(v) C_v]} &< \frac{\bar{Y}^2 C_u^2}{n} \\ - \frac{\rho^2 \bar{Y}^2 C_u^2}{n} - \frac{[\bar{X}_2 C_v \gamma_1(v) S_{uv} - \mu_{12}]^2}{n \bar{X}_2^4 [4C_v^2 + \beta_1(v) - 1 - 4\gamma_1(v) C_v]} &< 0 \end{aligned}$$

4.12 $\text{MSE}(\bar{Y}_{1pps}) < \text{MSE}(\bar{Y}_R)$

$$\begin{aligned} \frac{\bar{Y}^2 C_u^2 (1 - \rho^2)}{n} - \frac{[\bar{X}_2 C_v \gamma_1(v) S_{uv} - \mu_{12}]^2}{n \bar{X}_2^4 [4C_v^2 + \beta_1(v) - 1 - 4\gamma_1(v) C_v]} &< \frac{\bar{Y}^2}{n} [C_u^2 + C_v^2 - 2\rho C_u C_v] \\ - \frac{\bar{Y}^2 [\rho C_u - C_v]^2}{n} - \frac{[\bar{X}_2 C_v \gamma_1(v) S_{uv} - \mu_{12}]^2}{n \bar{X}_2^4 [4C_v^2 + \beta_1(v) - 1 - 4\gamma_1(v) C_v]} &< 0 \end{aligned}$$

4.13 $\text{MSE}(\bar{Y}_{1pps}) < \text{MSE}(\bar{Y}_p)$

$$\begin{aligned} \frac{\bar{Y}^2 C_u^2 (1 - \rho^2)}{n} - \frac{[\bar{X}_2 C_v \gamma_1(v) S_{uv} - \mu_{12}]^2}{n \bar{X}_2^4 [4C_v^2 + \beta_1(v) - 1 - 4\gamma_1(v) C_v]} &< \frac{\bar{Y}^2}{n} [C_u^2 + C_v^2 + 2\rho C_u C_v] \\ - \frac{\bar{Y}^2 [\rho C_u + C_v]^2}{n} - \frac{[\bar{X}_2 C_v \gamma_1(v) S_{uv} - \mu_{12}]^2}{n \bar{X}_2^4 [4C_v^2 + \beta_1(v) - 1 - 4\gamma_1(v) C_v]} &< 0 \end{aligned}$$

4.14 $\text{MSE}(\bar{Y}_{1pps}) < \text{MSE}(\bar{Y}_f)$

$$\begin{aligned} \frac{\bar{Y}^2 C_u^2 (1 - \rho^2)}{n} - \frac{[\bar{X}_2 C_v \gamma_1(v) S_{uv} - \mu_{12}]^2}{n \bar{X}_2^4 [4C_v^2 + \beta_1(v) - 1 - 4\gamma_1(v) C_v]} &< \frac{\bar{Y}^2 C_u^2 (1 - \rho^2)}{n} \\ - \frac{[\bar{X}_2 C_v \gamma_1(v) S_{uv} - \mu_{12}]^2}{n \bar{X}_2^4 [4C_v^2 + \beta_1(v) - 1 - 4\gamma_1(v) C_v]} &< 0 \end{aligned}$$

V. Empirical Study

To compare proposed Estimator numerically with other estimator. The description of population are given below
POPULATION (source: Singh D and Chaudhary F.S. Sample Survey Designs .ppt-132

$N=8, \bar{X}_1=36.25, \bar{X}_2=33.182, \bar{Y}=38.875$

$\bar{X}_2^2=1101.045124, \bar{X}_2^3=36534.879305, \bar{X}_2^4=1212300.3651$

$\bar{Y}^2=1511.265, \bar{Y}^3=58750.4269, \bar{Y}^4=2283922.85$

$\sigma_u^2=1.463, \sigma_v^2=392.222, C_u^2=0.00096, C_v^2=0.00556, \rho_{uv}=-0.2683$

$\gamma_1(v)=4394.23077, \beta_1(v)=564.919931, \mu_{12}=218.372735$

Mean Square Errors of $\bar{y}, \bar{y}_r, \bar{y}_p, \bar{y}_f, \bar{y}_{1pps}$

$M(\bar{y})=0.181351, M(\bar{y}_r)=1.462148, M(\bar{y}_p)=1.001213$

$M(\bar{y}_f)=0.168297, M(\bar{y}_{1pps})=0.163741$

Relative Efficiencies of $\bar{y}, \bar{y}_r, \bar{y}_p, \bar{y}_f$

$\bar{y}=110.754, \bar{y}_r=892.963$

$\bar{y}_p=611.461, \bar{y}_f=102.782$

VI. Conclusion

Thus the mean square error of our proposed estimator is less than the usual estimators Ratio, Product & Regression

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